Access to Credit and Rice Production Efficiency of Rural Households in the Mekong Delta

Vuong Quoc Duy,

School of Economics and Business Administration, CanTho University, Viet Nam. E-mail: vqduy@ctu.edu.vn

Abstract

Currently, rice production in the Mekong Delta region accounts for more than 50% of Vietnam's total paddy production and 90% of its rice export volume. Therefore, increasing the efficiency of rice production systems and enhancing the comparative advantage of Vietnam's rice industry have been an important focus area for policy makers and researchers for many years. Access to credit has been identified as a key factor for improving rice production. This fact is validated in this study by considering the production and technical efficiency levels of rice production for a sample of farmers in the Mekong Delta. The study focuses particularly on the effects of both formal and informal credits on production levels and production efficiency by using a Stochastic frontier analysis and a quantile regression. The results confirm the positive influence of credit on production and production efficiency. Both formal and informal credit appears to be important.

Key words: Stochastic frontier analysis, quantile regression, access to credit, rural

households

JEL Classification: E5, G2, O2

1. Introduction

Since Vietnam launched its innovation policy in 1986, the government has recognized the significant role played by agriculture. It has liberalized rice and agricultural input markets and implemented policies to promote the cultivation of high-yielding crop varieties. Since then Vietnam has experienced a steady increase in rice production and exports. Rice production reached 99 million tons in 2010 with rice yields of 5.32 tons/ha (GSO, 2010). The country has been a major rice exporter since 1989, currently among the top exporters in the world. In 2010, it exported as much as 6.88 million tons for 3.23 billion US dollars, up by 15.4% in volume and 21.2% from the previous year (GSO, 2010). These results were achieved with widely applied modern rice technology, resulting in the widespread adoption of high yielding modern rice varieties, whose use increased from 17% in 1980 to nearly 90% in 2000 (Ut and Kajisa, 2006, pp. 167-189).

The MD is recognized as the rice bowl of Vietnam and rice production is the main income source for farmers in the region. It accounts for more than 50% of the country's total paddy production and 90% of national rice export volumes (GSO, 2010). Yet not all farms seem to produce at their optimal level. There may possibly be room for improvement which could contribute to rural household income and continue to increase the comparative advantage of Vietnam's rice production. Credit access can be important in the further improvement of a farming system because of its ability to create access to other production factors (Oladeebo and Oladeebo, 2008, pp. 59-62). Rashid, et al. (2002) found that small farmers without credit in Bangladesh allocated less land to rice even when the magnitude of the effects of credit was very small. The availability of credit can affect fixed production costs (Brambilla and Porto, 2005) and farmers will use fewer seeds and fertilizer if they are credit constrained. Through its effect on production, output and covering marketing costs, credit may also influence farmers' participation in marketing systems and increase their ability to generate higher incomes.

This chapter investigates the contribution of credit to rice production and technical efficiency levels. Credit is different from input subsidies or technology provision because it is not free and loans need to be repaid at maturity (CGAP, 2006), which is why lenders require collateral. Previous studies on rice production have been concerned with the importance of credit on productivity, whether through Stochastic Frontier Analysis (SFA) or Data Envelopment Analysis (DEA) approaches. This chapter investigates whether scale matters for the effect of credit on production levels by applying a quantile regression, and a stochastic frontier to analyse the influence of credit access on efficiency levels is estimated. While descriptive analysis and the SFA have shown large differences in production and efficiency between borrowers and non-borrowers, it is not clear whether the positive effect is due to the

access to credit or to differences in scale. The quantile regression confirms that credit contributes to production across the different production quantiles. This chapter also distinguishes formal from informal credit in rice production functions.

The rest of this chapter is constructed as follows. Section 2 gives a literature review on studies on rice production efficiency conducted in Vietnam and elsewhere in the world. Section 3 presents the research methodology used in the present study. The empirical results of the study are given in section 4. Finally, section 5 offers a conclusion.

2. Literature Review

Technical efficiency in agricultural (including rice) production has been widely studied (Coelli and Battese, 1996, pp. 103-128, Kyi and Oppen, 1999, Coelli, 2002, pp. 607-626, Rahman, 2003, pp. 408-503, Amaza and Maurice, 2005, Moses and Adebayo, 2007, pp. 20-30). Coelli and Battese (1996, pp. 103-128) investigated the determinants of technical efficiency among Indian farmers using a two stage panel data approach consisting of a SFA and technical inefficiency effects models. Their findings showed that farmer age, level of education, farm size and survey years had a significant influence on the inefficiency effects of farmers in the study areas. Kyi and Oppen (1999) also used a SFA approach but paid more attention to physical and human capital in irrigated rice production in Burma. Their findings indicated that seed use, human resource capability and knowledge of farmers were significant determinants of rice productivity. An additional finding was that inefficiency effects seemed to be different for large farmers using fertilizer compared to small and large farmers not using fertilizer. Rahman (2003, pp. 408-503) analyzed profit efficiencies among Bangladeshi rice farmers using a stochastic profit frontier and an inefficiency effects model to calculate technical, allocative and scale efficiency levels. The results showed high levels of inefficiency. Variations in inefficiency levels were explained by differences in infrastructure, soil fertility, experience, extension services, tenancy and share of non-agricultural income.

In Nigeria, Amaza and Maurice (2005) investigated factors influencing the technical efficiency of rice-based production systems among farmers in Adamawa State. A SFA was applied on a cross sectional data set of 122 farmers collected in 2002-2003. The results showed that relative technical efficiency levels varied widely between farms, ranging between 0.26 and 0.97, with a mean technical efficiency of 0.80. They concluded that the efficiency of rice production among farmers in Adamawa State could increase by 20% through better use of available resources such as seeds, irrigation water and fertilizer, and through farmer education. Several other studies have been conducted in Nigeria using parametric or non-parametric analysis approaches (Idiong, 2007, pp. 653-658, Moses and Adebayo, 2007, pp. 20-30, Oladeebo and Fajuyigbe, 2007, pp. 93-100). Their findings suggest that human capital such as the farmer's educational level and age, as well as membership in a cooperative/farmer

association and access to credit has a significantly positive influence on efficiency (Idiong, 2007, pp. 653-658, Oladeebo and Fajuyigbe, 2007, pp. 93-100). In addition, physical capital, such as seed use, size of land and quantity of herbicide and/or fertilizer used (Moses and Adebayo, 2007, pp. 20-30) influences technical efficiency calculated for rice production.

In Vietnam, a number of studies have analysed rice production efficiency. Nguyen (2003, pp. 325-357) studied the technical efficiency in the MD using a SFA. Findings revealed that the average technical efficiencies in the three growing seasons - winter-spring, spring-summer and summer-autumn - are 86%, 79%, and 80%, respectively. Technical inefficiency was explained by variables such as land size, rice variety used, integrated pest management adoption, sowing techniques and availability of credit. Chi and Yamada (2005) also calculated the technical efficiency of rice farmers in Can Tho province using a production frontier function. Their results suggest that the technical efficiency of rice production is significantly affected by the education level and technical knowledge of the household head/farmer and the inputs used. Similarly, Vu (2007) study of the efficiency of Vietnamese rice farming households (based on both DEA and SFA) found that differences in technical efficiency could be explained by differences in length of education (as well as regional fixed effects).

In this chapter we break new ground by combining SFA and quantile regressions to analyze the contribution of credit to rice production efficiency and rice yields. And, as mentioned in the introduction, we suspect that scale economies may be involved. Some of the above-cited papers confirm that farm size matters for efficiency levels. This chapter applies a quantile regression to check whether scale of production also has an effect on the role played by access to credit.

3. Methodology

3.1 Research Questions

As indicated in introduction, we have two major research questions. One of them is to find out whether the access to credit effects on the rice efficiency of households in Mekong Delta. We will examine the volatility of the access to credit on both rice technical efficiency and scale of return on rice production of the households in Mekong Delta.

3.2 Modeling Volatility

3.2.1 Stochastic Frontier Analysis (SFA)

SFA, originally proposed by Aigner, et al. (1977, pp. 21-37) and Meeusen and Vandenbroeck (1977, pp. 435-445) and modified by Jondrow, et al. (1982, pp. 233-238), is popular in determining farm efficiency levels for cross-sectional data. It specifies the relationship between output and input levels using two error terms. One error term is the traditional normal error term in which the mean is zero and the variance is constant. The other

is the technical inefficiency level, expressed as a half normal, truncated normal, exponential or two parameter gamma distribution (Coelli, 1996a). The two error terms in the maximum likelihood estimation of the production function determine inefficiency levels.

In a SFA, output Y_i is function of inputs X_i as follows (Greene, 2008)

$$Y_i = f(x_i; \beta) + \varepsilon_i = f(x_i; \beta) + (v_i - u_i) \qquad i = 1....N$$
 (1)

Where: Y_i is the production level (or log production) of the i-th farm and X_i is a 1xK vector of transformation of input quantities of the i-th farm. The function f(.) is typically a Cobb-Douglas production technology or translog technology. Both functional forms are used extensively in the literature (Thiam, et al., 2001, pp. 235-243). The forms of Cobb-Douglas and Translog can be expressed as follows (Van Passel, et al., 2009, pp. 3057-3069):

Cobb-Douglas function:

$$\log(y_i) = \beta_0 + \sum_{k=1}^{n} \beta_k \log(x_{ik}) + v_i - u_i$$
 (2)

Translog function:

$$\log(y_i) = \beta_0 + \sum_{k=1}^n \beta_k \log(x_{ik}) + \sum_{k=1}^n \beta_{kk} (\log(x_{ik})^2 + \sum_{k=1}^n \beta_k \beta_t (\log(x_{ik}) \cdot \log(x_{ik})) + v_i - u_i$$
(3)

This chapter assumes a Cobb-Douglas model. A trans log specification would require a larger sample and may increase the probability of multicollinearity.

The error term in equation (1) is composed of two components (Aigner, et al., 1977, pp. 21-37):

$$\epsilon_i = v_i - u_i \,$$

where: v_i is the symmetric component which it accounts for random variation in output due to factors outside the farmer's control such as weather and diseases and which is assumed to be independently and identically distributed with $N(0, \delta_v^2)$;

 u_i captures technical inefficiency in production and is also assumed to be an independently and identically distributed non-negative truncation of the $N(\phi, \delta_u^2)$ distribution. In this chapter, we assume u_i to be half-normal distributed as specified by Greene (2008).

SFA also allows the estimation of the determinants of the TE levels in an inefficiency model. The variable u_i, which estimates technical inefficiency of household, is regressed against the household's socio-economic characteristics Z as follows (Coelli, 1996b):

$$u_i = \alpha_0 + \sum_{i=1}^n \alpha_i Z_i \tag{4}$$

The SF model enables us to estimate parameters and standard errors, and to test hypotheses by using the maximum likelihood method. The parameter vectors $\boldsymbol{\beta}$ and $\boldsymbol{\delta}$ are estimated together with the variance parameters $\delta^2 = \delta_v^2 + \delta_u^2$ and $\gamma = \delta_u^2 / \delta^2 = \frac{\delta_u^2}{\delta_v^2 + \delta_v^2}$.

The Frontier 4.1 program written by Coelli, et al. (1998) is used for the SFA estimation in this chapter .

3.2.2 Quantile regression

The Cobb-Douglas model estimated in the SFA does not allow allocation for non-linearity in production scale. To check scale effects on production and to check if the effect of credit differs over production scale, a quantile regression is estimated. The quantile regression model was first introduced by Koenker and Bassett (1978, pp. 33-50) and was developed in Koenker and Hallock (2001, pp. 143-156). The θ^{th} quantile for random variables X is defined as the value m_{θ} , such that the probability of X is less than m_{θ} . Mathematically this becomes (Koenker and Hallock, 2001, pp. 143-156):

$$\Theta = \Pr[X \le m_{\theta}] = F(m_{\theta}) \tag{5}$$

which is the cumulative distribution function of X. The value $m^{\wedge_{\theta}}$ for a sample quantile can be derived using the inverse of the cumulative distribution function, which is the quantile function Q(θ), under the assumption of a strictly monotonic, continuous distribution function, namely (Koenker and Hallock, 2001, pp. 143-156):

$$m^{\wedge} \theta = F^{-1}(\theta) = Q(\theta) = \inf \{ X \in \mathbb{R} : 0 \le F(X \theta) \}$$
 (6)

with inf $\{\}$ defined as the greatest lower bound of m^{\(\delta\)} θ . Therefore, the function Q(\theta) returns the lowest value for which the statement is true.

The term quantile model is used to distinguish between the numbers of equal size subsets used. For instance, four quantiles refer to quartile ($\theta = 0.25$; 0.5; 0.75 and 0.95) positions, dividing the data set into four equal-size groups. The 0.5 quantile is the median.

3.2.3 Model specification

Both the SFA and quantile models take rice production in tons (efficiency and yield) as the dependent variable. The independent variables considered for the Cobb-Douglas SFA model are the basic production factors: area under rice cultivation, expenditure on seed, fertilizer and pesticides, expenditure on hired labour and cost of hiring machines (table 1). Factors suspected to influence inefficiency levels (following the literature cited above) are household characteristics (age, gender, education level and experience of the head of household, family size and Vietnamese ethnicity), farm characteristics (use of new technology and access to amenities such as electricity), market orientation (percentage of sales and distance to the market centre), location and finally the uptake of formal or informal credit. The same variables are used in the quantile regression. The uptake of formal credit is

reflected by a dummy variable that takes value one if the farmer has borrowed a sum from one of the financial institutions in the MD, and zero otherwise. A similar dummy variable reflects the uptake of informal credit.

3.3 Data

The data used in this chapter is drawn from a household survey on living standards in Vietnam, the Vietnam Living Standard Survey – VLSS 2008. The survey was conducted by the General Statistical Office of Vietnam (GSO) in 2008, funded by United Nation Development Programme (UNDP) and Swedish International Development Cooperation Authority (SIDA) with technical assistance from the World Bank. It gathered information through community and household level questionnaires. The household questionnaires were organized into nine sections comprising basic demography, employment and labour force participation, education, health, income, expenditure, housing, fixed assets and durable goods, and participation in poverty reduction programmes.

The selection of the total sample of 45,945 households followed a method of stratified random cluster sampling of 3,063 communes/villages in Vietnam to make the data representative for national, rural and urban, and regional levels. The sample is divided into two subsamples: one sample with 9,189 households and another with 36,756 households (GSO, 2008). From the former sample, detailed information needed for household living standards analysis at national and regional levels were gathered; from the latter sample it was not.

As this chapter focuses on rice production of households in the Mekong Delta, a subsample of 654 households in this region was selected from the 9,189 households. The criteria for selection were residence in the MD, production of rice and availability of sufficiently detailed information on this production. Included in the data is information on household characteristics, access to amenities, market orientation and very importantly, rice input and output levels (table 1).

Table 1: Variable definition and measurement

Variables	Units	Definition
Rice production (Y)	Tons	Quantity of rice produced
Area rice (X_1)	hectare	Land area for rice planted
Seeds (X_2)	1,000 dongs/hectare	Expenditure on seed
Fertilizer (X_3)	1,000 dongs/hectare	Expenditure on chemical fertilizer
Pesticides (X_4)	1,000 dongs/hectare	Expenditure on pesticides
Hired labour (X_5)	1,000 dongs/hectare	Expenditure on labour hiring
Hired machinery (X ₆)	1,000 dongs/hectare	Expenditure on hiring farm implements
	-	for soil preparation and rice cutting

4. Results and Discussion

Table 2 compares the household and farm characteristics of borrowers and nonborrowers. There was no difference between both the groups in terms of age, farming experience and family size. The household heads of the borrower group have on average higher education levels. Rice production in the borrower group was significantly higher than that of non-borrowers. Borrowers planted more rice, and spent more on inputs for rice production such as fertilizers, pesticides and hired machinery than non-borrowers except for expenditure on seeds. Additionally, the household characteristics of borrowers and non-borrowers were not significantly different for any of the binary variables except the Vietnamese ethnicity and the adoption of new technology, where, not surprisingly, the percentage of adoption among borrowers was much higher (table 3).

Table 2: Household and farm characteristics of borrowers and non-borrowers

Independent variable	Non-borrowers	Borrowers	t- test	
n	312	342		
Age (years)	52.89	51.18	1.60	
	(0.799)	(0.727)		
Family size (persons)	4.46	4.55	-0.10	
	(0.097)	(0.085)		
Farming experience (years)	24.84	23.59	1.38	
	(0.679)	(0.608)		
Education (years)	5.64	6.90	-4.67***	
-	(0.186)	(0.194)	-4.0 / ****	
Rice production (tons)	21.68	42.633	-7.80***	
-	(1.262)	(2.294)	-/.80****	
Rice production (tons/ha)	5.11	5.584	-2.18*	
-	(1.377)	(3.672)		
Area with rice (ha)	4.83	7.943	5.16***	
	(8.225)	(7.261)	5.10****	
Seeds (1,000 dongs)	2,010	2,194	0.14	
-	(247.428)	(193.290)	-0.14	
Fertilizer (1,000dongs)	9,750	11,503	1 17	
-	(892.09)	(1,183)	-1.17	
Pesticides (1,000 dongs)	3,494	4,887	-1.86 **	
-	(387.763)	(623.019)	-1.80	
Hired labour (1,000 dongs)	3,990	3,310	0.720	
-	(896.371)	(387.573)		
Machinery hired (1,000 dongs)	9,733	13,223	1 02**	
	(1,115)	(1,513)	-1.83**	
Distance to market centre (m)	1,603	1,569	0.57	
	(42.668)	(40.113)	0.57	

Notes: *, **, ***: Significant at 10%; 5%; and 1%.

Table 3: Household characteristics of borrowers and non-borrowers

Independent variables	Non-Borrowers	Borrowers	X ² Test
n	312	342	
Gender (% male)	80.13	81.58	0.22
Vietnamese ethnicity (%)	93.07	90.26	2.18*
New technology (% yes)	16.35	93.86	399.70***
Access to electricity (% yes)	98.72	98.25	0.24
Can Tho province (% yes)	4.49	3.51	0.41

Notes: *, **, ***: Significant at 10%; 5%; and 1%.

The sources and characteristics of credit are presented in table 4. Most borrowers (68.1%) had taken out a loan from a formal financial institution (Vietnam Bank of Agriculture and Rural Development or Vietnam Bank Social Policy). About 31.4% of the borrowers had a

loan from informal sources who were individual lenders, friends and relatives. The remaining 1.5% had a loan from semi-formal lenders such as job creation funds and social political associations. There was a difference between loan characteristics among providers. Average loans taken out from formal lenders were larger than those of the other credit sources. The informal lenders charged higher interest rates.

Table 4: Loan sources and characteristics

Loan	Units	Formal	Semi-formal	Informal	F-Test
characteristics/Source		sector		sector	
Loan size	1000 dongs	29,635	10,214	19,844	2.27***
		(49,596)	(6,903)	(28,047)	
Interest rate	%/year	16	8.40	26.00	12.74***
	-	(7.211)	(2.003)	(31.941)	
Maturity	months	16	19	14	2.67**
-		(12.778)	(17.904)	(9.712)	
Borrowers in sample	n	233	5	104	342
Percentage	%	68.1	1.5	31.4	100

Notes: Standard deviation in parentheses; *, **, ***: Significant at 10%; 5%; and 1%.

As noted above, the production function was analysed for four quartiles of rice production (table 5). Quartile 95 dominates the other quartiles with respect to all variables that were included in the model. Age and education were lowest for the households in quartile 25. Households in the higher quartiles were using more inputs. In particular, those in quartile 95 were spending more on inputs than the other quartiles by a factor of 3 to 30 depending on the type of inputs.

Table 5: Household characteristics of household quartiles

Independent variable	Mean	Standard Deviation	Q25	Q50	Q75	Q95
Rice production (tons)	32.634	35.869	11.760	23.899	42.000	86.220
Area with rice (ha)	6.453	7.886	2.970	4.680	7.500	17.520
Rice per ha (tons)	5.057	4.548	3.960	5.107	5.60	4.921
Expenditure (1000 dongs/ha)						
Seeds	2,106	3,972	523	1,165	2,361	6,490
Fertilizer	10,665	19,219	2,315	5,400	11,234	39,802
Pesticides	4,222	9,599	595	1,603	4,284	16,032
Hired labour	3,635	12,096	500	1,612	3,904	10,738
Machines hired	11,558	24,433	2,971	5,881	11,969	34,401
Age (years)	51.990	13.780	42	50	61	77
Education (years)	5.150	3.110	3	5	8	12
Family size (persons)	4.510	1.640	3	4	5	7
Farming experience (years)	24.180	11.630	15	23	33	43

Table 6 gives the results of the maximum likelihood estimates (MLE) of the pooled household model with dummy variables for the household who took formal and informal credit. The average technical efficiency of rice production was about 85%. This implies that farmers still had the possibility to improve their efficiency by an average of 15%. Similar results were found by Awotide and Adejobi (2006, pp. 107-113) and Nguyen (2003, pp. 325-357), although the calculated efficiency levels are relative to the frontier of the sampled

households. In the Cobb-Douglas function, the coefficients of areas with rice and expenditure on pesticides were statistically significant.

Table 6: Estimation of SFA and efficiency levels for rice production

14010 01 201111		SFA model	SFA model				
Production Coefficient	Units	Co-efficient	Standard Deviation	t-ratio			
Dependent variable:				_			
Log(rice outcome) (tons)							
Independent variables (Log):							
Constant	β0	1.294***	0.181	7.135			
Area rice (ha)	β1	0.741***	0.062	11.881			
Seeds (1,000 dongs)	β2	-0.073	0.050	-1.452			
Fertilizer (1,000 dongs)	β3	-0.033	0.046	-0.723			
Pesticides (1,000 dongs)	β4	0.063**	0.032	1.965			
Hired labour (1,000 dongs)	β5	0.033	0.028	1.155			
Hired machines (1,000 dongs)	β6	-0.019	0.043	-0.454			
δ2		0.048	0.007	7.259			
Γ		0.079	0.154	0.513			
Log Likelihood		61.856					
Inefficiency Effect Model							
Age (years)	δ1	-0.023	0.114	-0.204			
Family-size (persons)	δ2	-0.008	0.096	-0.079			
Experiences (years)	δ3	0.054	0.069	0.764			
Education (years)	δ4	-0.109*	0.061	-1.778			
Vietnamese ethnicity (yes=1)	δ5	0.055	0.063	0.879			
Gender (male=1)	δ6	0.015	0.041	0.357			
New technology (yes=1)	δ7	-0.129**	0.047	-2.757			
Formal credit (yes=1)	δ8	-0.474**	0.182	-2.605			
Informal credit (yes=1)	δ9	-0.299**	0.115	-2.613			
Can Tho province (yes=1)	δ10	-0.190**	0.076	-2.501			
Distance to market (1,000 m)	δ11	0.112*	0.065	1.738			
δυ2		0.004		_			
δν2		0.037					
TE		0.851					
LR test of the one sided error		302.571					

Notes: *, **, ***: Significant at 10%; 5%; and 1%.

The model indicated that technical inefficiency of rice production was associated with household characteristics (education level), technology (use of new farm technology), location (distance from households' dwelling place to market centre), and the Can Tho province and the access to formal or informal credit. This implies that rice efficiency was positively associated with better educated households using newer techniques, households from the Can Tho province and those further out from the market centre. Both formal and informal credit seems to increase farm efficiency. The coefficients of average age, family size, farming experience, Vietnamese ethnicity and gender of household head were not significant. The likelihood ratio tests of one-side generalized error exceeded the critical value (α =5 percent), suggesting that the null hypothesis of no technical inefficiency in rice production in the sample is rejected.

The distribution of technical efficiency of rural households in the MD is presented in table 7. The technical efficiency levels of borrowers were higher than those of non-borrowers.

Table 7: Distribution technical efficiencies of households

D	Borrowe	Borrowers				Non-borrowers	
Range	Formal borrowers		Informal borrowers				
	n	%	n	%	n	%	
<70%	0	0	0	0	157	50	
70-80%	0	0	0	0	128	41	
80-90%	0	0	5	5	27	9	
90-100%	233	100	99	95	0	0	
Total	233	100	104	100	312	100	

Table 8 gives the results of the quantile analysis. The rice yields of all quantiles are significantly positively affected by the area with rice, expenditure on pesticides (except quantile 95), education level of the household head (except quantile 95) and use of new farming technology. Furthermore, for all quantiles, access to credit, both formal and informal, contributed significantly to production. The rice yield of quantile 50 was positively affected by age of the household head, Vietnamese ethnicity, higher expenditure on hired labour and location in Can Tho province. The coefficients of the dummy variables for access to formal and informal credit confirm that borrowers were more likely to have higher rice production outcomes than non-borrowers. The coefficients of formal credit uptake were larger than those for informal credit access in functions of quantile 25 and 95, and smaller in functions of quantile 50 and 75, but the differences were not large.

Table 8: Quantile regressions of rice production

	Units	Q25	Q50	Q75	Q95
Dependent variable:					
Log(rice outcome) (tons)					
Independent variables (Log):					
Constants	β0	0.221	0.280	1.027***	1.048*
		(0.97)	(1.12)	(3.44)	(1.73)
Area rice (ha)	β1	0.760***	0.698***	0.718***	0.803***
		(12.83)	(11.30)	(10.35)	(5.65)
Seeds (1,000 dongs)	β2	-0.058	-0.025	-0.042	-0.146
		(-1.20)	(-0.46)	(-0.66)	(-1.17)
Fertilizer (1,000 dongs)	β3	-0.009	0.060	-0.036	-0.095
		(-0.19)	(1.21)	(-0.65)	(-1.02)
Pesticides (1,000 dongs)	β4	0.128***	0.103***	0.080**	-0.020
		(4.34)	(3.10)	(2.03)	(-0.24)
Hired labour (1,000 dongs)	β5	0.009	-0.001	0.029	0.029
		(0.37)	(-0.00)	(0.84)	(0.41)
Hired machines (1,000 dongs)	β6	-0.027	-0.079*	-0.091	0.052
		(-0.71)	(-1.71)	(-1.64)	(0.41)
Age (years)	δ1	0.131*	0.173*	0.104	0.186
		(1.81)	(2.01)	(0.98)	(0.71)
Family-size (persons)	δ2	-0.025	-0.064	-0.071	-0.033
		(-0.47)	(-1.18)	(-1.13)	(-0.24)
Experience (years)	δ3	-0.049	-0.018	-0.038	-0.028
		(-1.40)	(-0.47)	(-0.84)	(-0.30)
Education (years)	δ4	0.099***	0.106**	0.157***	0.141
		(3.51)	(3.17)	(3.40)	(1.25)
Vietnamese ethnicity (yes=1)	δ5	0.048	0.056*	-0.085*	-0.057

Proceedings of the Second Asia-Pacific Conference on Global Business, Economics, Finance and Social Sciences (AP15Vietnam Conference) ISBN: 978-1-63415-833-6

Danang, Vietnam, 10-12 July 2015 Paper ID: V583

		(1.58)	(1.68)	(-2.22)	(-0.73)
Gender (male=1)	δ6	0.016	0.004	-0.038	-0.034
		(0.74)	(0.16)	(-1.31)	(-0.60)
New technology (yes=1)	δ7	0.101***	0.092***	0.086*	0.200***
		(3.66)	(3.07)	(2.24)	(3.29)
Formal credit (yes=1)	δ8	0.111***	0.189***	0.249***	0.334***
		(3.94)	(6.23)	(6.69)	(7.01)
Informal credit (yes=1)	δ9	0.097**	0.223***	0.319***	0.274***
		(2.99)	(6.44)	(7.73)	(4.34)
Can Tho province (yes=1)	δ10	0.079*	0.076*	0.0807	0.025
		(1.85)	(1.66)	(1.45)	(0.21)
Distance to market (1,000 m)	δ11	0.014	0.067	0.068	0.019
		(0.32)	(1.41)	(1.24)	(0.16)
N		654	654	654	654
Pseudo R-square		0.605	0.526	0.481	0.473
Min sum of deviations		78.972	105.675	87.031	28.463

Notes: t statistics in parentheses; *, **, ***: Significant at 10%; 5%; and 1%.

The possible explanations for the model results are as follows. First, the coefficients of access to credit were negatively significant in technical inefficiency regression and positively significant in the quantile regression, implying that access to credit is likely to increase the production and technical efficiency of rice farmers. In fact, the financial constraints in farming are likely to be relieved through credit, which allows the purchase of more inputs, which in turn increases revenues and profit (Hyuha, et al., 2007, pp. 243-253). Additional funds from credit markets can be used to invest in rice paddy production, principally by adopting new technologies (Nuryartono., et al., 2005). Nuryartono (2005) also found that access to financial markets facilitated the adoption of technology such as fertilizer and pesticides. In the inefficiency model, the coefficient of access to formal credit is larger than that of informal credit, suggesting that access to formal credit had a larger effect on rice production efficiency. The findings are consistent with Kebede (2001), Nwaru (2001, pp. 1-10), Ajibefund and Aderinola (2003), Nguyen (2003, pp. 325-357), Ogundari (2008, pp. 224-233).

A second possible explanation is that the coefficients for education and use of new technique were negative and significant in the technical inefficiency model and positively significant in quantile regression (except in the largest quantile). Education may enhance the acquisition and utilization of information on improved technology as well as their entrepreneurship (Coelli and Battese, 1996, pp. 103-128, Dey, et al., 2000, pp. 33-46, Effiong, 2005, Onyenweaku, et al., 2005, pp. 20-25, Idiong, 2006). The importance of the introduction of new technologies on production is also confirmed.

Third, the signs in both models of the coefficient of distance from household dwelling to the market suggest that households in remote areas were more likely to have lower technical efficiency levels and rice yields. Poor communication and transportation facilities may lead to lower efficiency levels of households further away from market centre. This finding is consistent with those in studies by Lanzona and Evenson (1997), DeSilva, et al. (2006, pp. 851-865), Larson and Plessmann (2009, pp. 24-32).

Furthermore, farmers in Can Tho province had higher production and efficiency levels than those in the other provinces. Rice farming in this province benefits not only from the natural fertilization provided by the Tien and Hau rivers, but also from the support of the two largest agricultural science centres of the Mekong Delta, namely Can Tho University and Cuu Long Delta Rice Research Institute.

5. Conclusions and Recommendations

This chapter investigates the technical efficiencies and yields of rice farmers in the MD of Vietnam by using SFA and quantile regressions. The determinants of the stochastic production frontier tested included the land area used for rice and expenditure on seeds, hired labour, fertilizer, pesticides and hired machines. Coefficients of area with rice and of expenditure on pesticides had the expected signs (as they did in studies by Coelli and Battese (1996, pp. 103-128), Kyi and Oppen (1999), Wadud and White (2000, pp. 1665-1673), Jaforullah and Premachandra (2003), Nguyen (2003, pp. 325-357), Ogundari (2008, pp. 224-233)). Technical efficiency and rice yields were positively influenced by access to credit, household characteristics (educational level of household head), location of households (location in Can Tho province and proximity to the nearest market centre), farm technology (use of new farm technology and expenditure on pesticides) and area with rice. In addition, borrowers are relatively wealthier than non-borrowers, although the quantile regression confirmed that credit positively contributes to production among the smaller producers as well.

The results of this study have a number of implications, especially for credit accessibility. They have shown that access to formal credit had a larger effect on rice production efficiency than the uptake of informal credit. Formal credit is regulated while the informal credit is not and is easier to access. A further expansion of rural credit systems could enhance and contribute to increased rice production and efficiency in the Mekong Delta. Given the limited extent of governmental credit programmes in the Mekong Delta, accessibility to credit by rural households could be improved by establishing more branches of agricultural and community banks in the rural areas, providing innovative credit schemes that overcome problems of smallholder farmers who lack collateral and by reducing the currently long processing times of loan applications and other requirements. In addition, access to credit could be made easier for farmers without the specific commodity requirements. Furthermore, credit should meet the needs of the farmers, in particular for investment in farm activities.

Credit awareness and the establishment of strong and viable farmer organisations (such as cooperatives or credit associations) which can play a leading role in increasing farmers'

access to credit are important. Similarly, savings mobilization programmes should be developed and promoted in the survey area, which will inspire participation and provide encouragement to farmers to save and reinvest. Savings programmes also reduce the costs of monitoring lenders.

Many rural clients of the formal credit programmes have a lack of skills and training, and limited access to markets and technology. As a result, when these households have access to credit to invest in an existing business or to start-up a new one, the sustainability of their activities may become problematic. Therefore, it is very important for the financial institutions to facilitate or directly involve themselves in "credit plus" services that may include skill development/training, marketing facilities and business development services for their customers to help them sustain the economic activities supported by their financial schemes. To a more general extent, education is needed to improve efficiency levels. This could also increase the use of new farming technology, which plays a significant role in rice production. A future research project should focus on the impact of external services intervention on rice farming by panel data.

References

Aigner, D. J., Lovell, C. A. K., Schmidt, P., 1977. Formulation and estimation of stochastic frontier production function models, *Journal Economics* 6, 21-37.

Ajibefund, I. A., Aderinola, A., 2003. Determinants of technical efficiency and policy implications in traditional agricultural production: empirical study of Nigerian food crop farmers, Work in progress report presented at the bi-Annual research Workshop of AERC, May 24-29th, Nairobi Kenya.

Awotide, D. O., Adejobi, A. O., 2006. Technical Efficiency and Cost of production of Plantain farmers in Oyo State, Nigeria, *Moor Journal of Agricultural Research*. 7, 107-113.

Brambilla, I., Porto, G. G., 2005. Farm Productivity and Market Structure: Evidence From Cotton Reforms in Zambia, Access on 29/03/2012 at:

http://www.depeco.econo.unlp.edu.ar/semi/semi030605.pdf

CGAP, 2006. Commercial Loan Agreements A Technical Guide for Microfinance Institutions, Consultative Group to Assist the Poor Prepared for CGAP by Cleary, Gottlieb, Steen & Hamilton, LLP.

Chi, T. T. N., Yamada, R., 2005. Assessing the technical efficiency of input rice production in Thoi Lai Commune, Co Do district, Can Tho City, Omonrice.

Coelli, T. J., 1996b. Guide to FRONTIER Version 4.1: A Computer Program for Stochastic Frontier Production and Cost Function Estimation, CEPA Working Paper 96/7. Armidale NSW Australia., Department of Econometrics, University of New England.

Coelli, T., 1996a. A Guide to FRONTIER Version 4.1: A Computer Program for Stochastic Frontier Production and Cost Function Estimation. Centre for Efficiency and Productivity Analysis, University of New England, Armidale, NSW, 2351, CEPA Working Paper 96/07.

Coelli, T., Rao, D. S. P., Battese, G. E., 1998. An introduction to efficiency and productivity analysis, , Kluwer Academic Publishers, Boston/Dordrecht/London.

DeSilva, S., Evenson., R. E., Kimhi., A., 2006. Labor Supervision and Institutional Conditions: Evidence from Bicol Rice Farms, *American Journal Agricultural Economics* 88, 851-865.

Dey, M. M., Paraguas, F. J., Bimbaa, G. B., Ragaspi, P. B., 2000. Technical efficiency of tilapia growth out pond operations in the Philippines, *Agricultural Economics and Management 4*, 33-46.

Effiong, E. O., 2005. Efficiency of production in selected livestock enterprises in Akwa Ibom State, Nigeria., Unpublished PhD Dissertation. Michael Okpara University of Agriculture, Umudike.

Greene, W. H., 2008. Econometric Analysis, in 6th ed. ed. New Jersey: Prentice Hall.

GSO, 2010. General Statistics Office: Statistical Yearbook of Vietnam 2009 (in Vietnamese), *Statistical Publishing House, Hanoi.*

Hyuha, T. S., Bashaasha, B., Nkonya, E., Kraybill, D., 2007. Analysis Of Profit Inefficiency In Rice Production in Eastern And Northern Uganda, *African Crop Science Journal* 15, 243-253.

Idiong, I. C., 2006. Evaluation of technical, allocative, and economic efficiencies in rice production systems in cross river state, Nigeria., Unpublished Ph.D Dissertation. Michael Okpara University of Agriculture, Umudike.

Idiong, I. C., 2007. Estimation of Farm Level Technical Efficiency in Small Scale Swamp Rice Production in Cross River State of Nigeria: A Stochastic Frontier Approach, *World Journal of Agricultural Sciences 3*, 653-658.

Jaforullah, M., Premachandra, E., 2003. Sensitivity of technical efficiency estimates of estimation approaches: An investigation using New Zealand dairy University of Otago, Economics Discussions Papers 0306. New Zealand.

Kebede, T. A., 2001. Farm household technical efficiency: a stochastic frontier analysis a study of rice producers in Mardi-Watershed in the Western Development Region of Nepal, Master Thesis, Agricultural University of Norway.

Koenker, R., Bassett, G., 1978. Regression Quantiles, Economitrica. 46, 33-50.

Koenker, R., Hallock, K. F., 2001. Quantile Regression, *Journal of Economic Perspectives* 15, 143-156.

Kyi, T., Oppen, M. V., 1999. Stochastic Frontier Production Function and Technical Efficiency Estimation: A case study on irrigated rice in Myanmar, Deutscher Tropentag 1999 in Berlin, Session: Sustainable Technology Development in Crop Production.

Lanzona, L. A., Evenson, R. E., 1997. The Effects of Transaction Costs on Labor Market Participation and Earnings: Evidence from Rural Philippine Markets, Center Discussion Paper No. 790. Economic Growth Center: Yale University.

Larson, D. F., Plessmann, F., 2009. Do Farmers Choose To Be Inefficient? Evidence from Bicol, *Journal Development Economics* 90, 24-32.

Meeusen, W., Vandenbroeck, J., 1977. Efficiency Estimation From Cobb-Douglas Production Functions With Composed Error, *International Economics Reviews* 18, 435-445.

Moser, S. C., 2005. Impact assessment and policy response to sea level rise in three US states: An exploration of human dimention uncertainties, *Global Environmental Change 15*.

Moses, J., Adebayo, A. F., 2007. Efficiency of Factors Determining Rainfed Rice Production in Ganye Local Government Area, Adamawa State, *Journal of Sustainable Development in Agriculture and Environment.* 3, 20-30.

Nguyen, T. M. H., 2003. A Study on Technical Efficiency of Rice Production in The Mekong Delta-Vietnam by Stochastic Frontier Analysis, *Journal of the Faculty of Agriculture Kyushu University*. 48, 325-357.

Nguyen, T. M. H., 2003. A Study on Technical Efficiency of Rice Production in The Mekong Delta-Vietnam by Stochastic Frontier Analysis, *Journal of the Faculty of Agriculture Kyushu University* 48, 325-357.

Nuryartono, N., 2005. Impact of Smallholders' access to Land and Credit Markets on Technology Adoption and Land Uses Decision: The Case of Tropical Forest Margins in Central Sulawesi Indonesia. PhD Thesis, Insitute of Rural Development, George August University of Goettingen.

Nuryartono., N., Zeller., M., Schwarze., S., 2005. Credit Rationing of Farm Households and Agricultural production: Empirical Evidence in the Rural Areas of Central Sulawesi, Indonesia. Institut of Rural Development. Georg August University of Goettingen, Tropentag 2005 Stuttgart-Hohenheim, October 11-13, 2005

Nwaru, J. C., 2001. Gender and Relative Production Efficiency in Food Crop Farming in Abia State of Nigeria, *The Nigerian Agricultural Journal.* 34, 1-10.

Ogundari, K., 2008. Resource-productivity, allocative efficiency and determinants of technical efficiency of rainfed rice farmers: A guide for food security policy in Nigeria, *Agricultural Economics* 54, 224-233.

Oladeebo, J. O., Fajuyigbe, A. A., 2007. Technical Efficiency of Men and Women Upland Rice Farmers in Osun State, Nigeria, *Journal of Human Ecology* 22, 93-100.

Oladeebo, J. O., Oladeebo, O. E., 2008. Determinants of Loan Repayment among Smallholder Farmers in Ogbomoso Agricultural Zone of Oyo State, Nigeria, *Journal Social Science* 17, 59-62.

Onyenweaku, C. E., Igwe, K. C., Mbanasor, J. A., 2005. Application of a stochastic frontier production function to the measurement of technical efficiency in yam production in Nasarawa State, Nigeria, *Journal of Sustainable Tropical Agricultural Research* 13, 20-25.

Rashid, S., Sharma, M., Zeller, M., 2002. Micro-lending for small farmers in Bangladesh: Does it affect farm households' land allocation decisions?, *MSSD Discussion Paper 45*, International Food Policy Research Institute.

Thiam, A., Bravo-Ureta, B. E., Rivas, T. E., 2001. Technical efficiency in developing country agriculture: A meta-analysis, *Agricultural Economics* 25, 235-243.

Ut, T. T., Kajisa, K., 2006. The impact of Green Revolution on rice production in Vietnam, *Development Economics* 44, 167-189.

Van Passel, S., Van Huylenbroeck, G., Lauwers, L., Mathijs, E., 2009. Sustainable value assessment of farms using frontier efficiency benchmarks, *Journal Environmental Managment* 90, 3057-3069.

Vu, L. H., 2007. Efficiency of Rice Farming Households in Vietnam: A DEA with Bootstrap and Stochastic Frontier Application, Mimeo, Department of Applied Economics, University of Minnesota.

Wadud, A., White, B., 2000. Farm household efficiency in Bangladesh: A comparison of stochastic frontier and DEA methods, *Applied Economics* 32, 1665-1673.